

# Modelling of Dependent Credit Rating Transitions

Verena Goldammer

(Joint work with Uwe Schmock)

Financial and Actuarial Mathematics  
Vienna University of Technology

Wien, 15.07.2010

## Motivation:

- Volcano on Iceland erupted and caused that most of the flights in Europe had to be cancelled for a few days.
- That caused simultaneous losses of the airlines.

⇒ Credit quality of the airlines is simultaneously affected.

## Previous literature:

- Dependence introduced by interacting intensities
- No simultaneous credit rating transitions possible!

## Main modeling assumption:

Firms may simultaneously change their credit rating in continuous time.

## 1 Model

- General framework
- General model
- Examples

## 2 Simulation

## 3 Maximum Likelihood Estimation

- MLE for the extended strongly coupled random walk
- Asymptotic properties of the estimator

## Definition (Marked point process)

- $(\tau_i)_{i \in \mathbb{N}}$ : random time with values in  $(0, \infty]$ , and  $\tau_i < \tau_{i+1}$  on  $\{\tau_i < \infty\}$  and  $\tau_i = \tau_{i+1} = \infty$  on  $\{\tau_i = \infty\}$
- $(\rho_i)_{i \in \mathbb{N}}$ : random mark with  $\rho_i \in E$  on  $\{\tau_i < \infty\}$  and  $\rho_i := \rho_\infty$  on  $\{\tau_i = \infty\}$ , where  $\rho_\infty$  external point of  $E$ .

We call  $((\tau_i, \rho_i))_{i \in \mathbb{N}}$  a **marked point process**.

## Mark space $E$ :

$$E = \{r : S \times I \rightarrow S \mid r \text{ is } (\mathcal{P}(S) \otimes \mathcal{I})\text{-}\mathcal{P}(S) \text{ measurable}\}$$

- $S = \{1, \dots, K\}$ : credit rating classes, where  $K$  means firm is in default and 1 is best rating class
- Measurable space  $(I, \mathcal{I})$ :  
state space of idiosyncratic component

### Model

General  
framework  
General model  
Examples

### Simulation

Likelihood  
Estimation

# The General Framework

- $F = \{1, \dots, n\}$ : set of firms,  $n \in \mathbb{N}$  is number of firms
- $X = ((X_t(1), \dots, X_t(n)))_{t \geq 0}$ : credit rating process
- $((\tau_i, \rho_i))_{i \in \mathbb{N}}$ : marked point process
- $U_i(j)$ :  $I$ -valued random variable for  $i \in \mathbb{N}$  and  $j \in F$ .

## Model

General  
framework  
General model  
Examples

## Simulation

Likelihood  
Estimation

### Definition (General framework)

We say that the process  $X = (X_t)_{t \geq 0}$  with state space  $S^n$  **follows the general framework**, if

- 1  $X_t = X_0$  for  $t \in [0, \tau_1)$ , and
- 2 for each  $i \in \mathbb{N}$  and firm  $j \in F$

$$X_t(j) = \rho_i(X_{\tau_i-}(j), U_i(j)) \quad \text{for } t \in [\tau_i, \tau_{i+1}).$$

**Remark:** Process is in general not Markovian.

## Additional assumptions to obtain a Markov process:

- 1 Random times  $(\tau_i)_{i \in \mathbb{N}}$ :  
jump times of a Poisson process with intensity  $\lambda > 0$
- 2 Random marks  $(\rho_i)_{i \in \mathbb{N}}$ : i. i. d. sequence
- 3 Idiosyncratic components  $\{U_i(j) : i \in \mathbb{N}, j \in F\}$ :  
i. i. d. collection
- 4  $(\rho_i)_{i \in \mathbb{N}}$ ,  $\{U_i(j) : i \in \mathbb{N}, j \in F\}$ ,  $X_0$  and the Poisson process are pairwise independent.

In the following:

We assume that these additional assumptions are satisfied.

## Assumption for the general model:

All firms with the same rating may simultaneously change only to the same rating class or remain in their rating class.

## Dynamics of the general model:

- Possible rating transitions are given by a map  $s : S \rightarrow S$ :
  - Each firm with rating 1 either remains in this class or changes its rating to  $s(1)$ ,
  - each firm with rating 2 remains in 2 or changes to  $s(2)$ ,
  - and so on ...
- The probability that a firm actually changes is given by  $p_x$ , where  $x \in S$  is the current rating of the firm.

## Definition (General model)

We say that the Markov jump process  $X = (X_t)_{t \geq 0}$  **follows the general model with parameters**  $(\lambda, P, \rho)$ , if it follows the general framework with the additional assumptions:

- $P$  probability distribution on  $S^S$  and  $\rho \in [0, 1]^S$
- Each  $\rho_i$  takes a.s. only values in  $\{r_s : s \in S^S\} \subset E$  where

$$r_s(x, u) = \begin{cases} s(x), & \text{if } u \in [0, \rho_x], \\ x, & \text{if } u \in [\rho_x, 1]. \end{cases}$$

- $\mathbf{P}[\rho_i = r_s] = P(s)$  for each  $s \in S$
- $U_i(j)$ : uniformly distributed on  $I = [0, 1]$  for  $i \in \mathbb{N}$ ,  $j \in F$



# Example 1: The Strongly Coupled Random Walk

## Dynamics:

Only firms in one rating class may simultaneously change to the same rating class or remain in their rating class.

## Parameters:

- Independent Poisson processes with intensity  $\lambda_x > 0$  for each rating class  $x \in S$
- Stochastic transition function  $P^c : S \times S \rightarrow [0, 1]$ : probability for transitions from  $x$  to  $y$  given Poisson process of  $x$  jumps
- $p_x \in [0, 1]$ : probability that a firm with rating  $x$  actually changes the rating

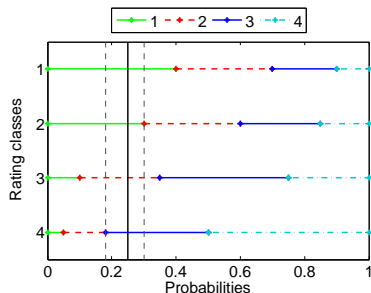
Define  $\lambda = \sum_{x \in S} \lambda_x$  and the distribution  $P$  on  $S^S$  by

$$P(s) = \begin{cases} \frac{\lambda_x}{\lambda} P^c(x, y), & \text{if there exist } x, y \in S \text{ with } x \neq y, \\ & s(x) = y, s(u) = u \text{ for all } u \in S \setminus \{x\} \\ \sum_{x \in S} \frac{\lambda_x}{\lambda} P^c(x, x), & \text{if } s(x) = x \text{ for all } x \in S, \\ 0, & \text{otherwise.} \end{cases}$$

## Definition (Strongly coupled random walk)

We say that the Markov jump process  $X$  is a **strongly coupled random walk process** with parameters  $((\lambda_x)_{x \in S}, P^c, p)$ , if  $X$  follows the general model with parameters  $(\lambda, P, p)$ .

## Example 2: The Scheme Model



For each  $x \in S$  the interval  $[0, 1]$  is divided into  $K$  subintervals with length  $p_{xy}$  for the  $y$ -th subinterval.

The subinterval containing  $V$  represents the rating class  $\tilde{s}(x)$ .

- $(p_{xy})_{x,y \in S} \in [0, 1]^{S \times S}$ : stochastic transition function
- $V$ : random variable, uniformly distributed on  $[0, 1]$
- $S^S$ -valued random function  $\tilde{s}$ :

$$\tilde{s}(x) = \max \left\{ y \in S : \sum_{k=1}^{y-1} p_{xk} \leq V \right\}, \quad \text{for } x \in S.$$

# Definition of the Scheme Model

The distribution of  $\tilde{s}$  is given by

$$P^S(s) = \max \left\{ \min_{x \in S} \sum_{k=1}^{s(x)} p_{xk} - \max_{x \in S} \sum_{k=1}^{s(x)-1} p_{xk}, 0 \right\}.$$

## Definition (Scheme model)

- $(p_{xy})_{x,y \in S} \in [0, 1]^{S \times S}$ : stochastic transition function
- $P^S$ : probability distribution of  $\tilde{s}$
- $\lambda > 0$  and  $p = (p_x)_{x \in S}$  is a vector in  $[0, 1]^S$

We say that the Markov jump process  $X$  follows the **scheme model** with parameters  $(\lambda, (p_{xy})_{x,y \in S}, p)$ , if  $X$  follows the general model with parameters  $(\lambda, P^S, p)$ .

Model

General  
framework  
General model  
Examples

Simulation

Likelihood  
Estimation

## Theorem (Embedding property)

- $X$  rating process in general framework with  $n$  firms
- $Y$  rating process in general framework with  $m < n$  firms

$\Rightarrow$  *Distribution of rating transitions of first  $m$  firms of  $X$   
= Distribution of rating transitions of  $Y$*

- Q-matrix  $\mu \in \mathbb{R}^{K \times K}$  of the transitions of the individual firms is the same for all firms.
- Correspondence of parameters:

$$(\mu, \rho) \Rightarrow (\lambda_x, P^c, \rho) \quad \text{or} \quad (\lambda, P^s, \rho)$$

- **Extended strongly coupled random walk:**  
 $\rho_x = 0$ : independent rating transitions of firms in class  $x$

# Loss of a Credit Portfolio

## Credit portfolio:

- $n = 100$  credits with amount  $C = 1$  and maturity  $T = 15$ .
- Obligors change credit rating according to process  $X$ .
- $K = 8$  rating classes, default  $K$  is an absorbing state.
- Recovery rate:  $\delta = 0.4$
- Default-free interest rate is zero.

## Loss of the credit portfolio:

$$L(t) = \sum_{i=1}^n C (1 - \delta) \mathbb{1}_{\{X_{t \wedge T}(i) = K\}}, \quad \text{for } t \geq 0.$$

# Empirical Excess Loss Distribution

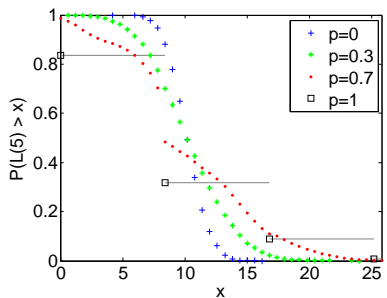
- $X_0$ : 16 firms in rating class 1, 14 firms each in 2 to 7
- $p_x = p$  for all  $x \in S$
- Intensity  $\mu$  of individual credit rating transitions is based on data of Standard & Poor's.

Model

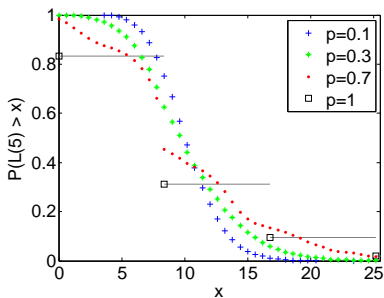
Simulation

Likelihood  
Estimation

**Empirical excess loss distribution (5 000 simulations):**

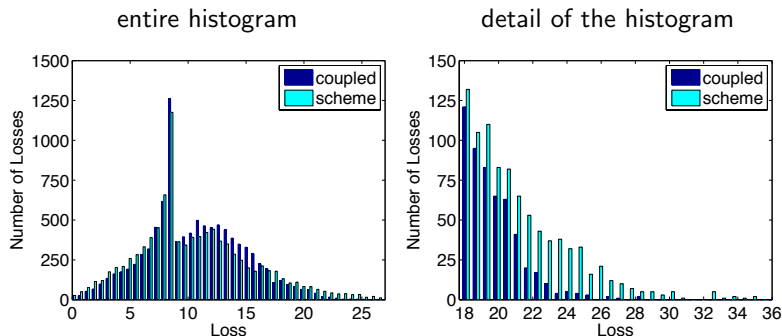


strongly coupled random walk



scheme model

# Histogram of Simulated Losses



**Figure:** Histogram of the simulated losses for the strongly coupled random walk (coupled) and the scheme model (scheme) where  $\rho = 0.5$  and  $t = 5$ , based on 10 000 simulations.



## Notation:

- Set of parameters:  $\Theta = ([0, \infty)^{K-1} \times [0, 1])^K$
- $\theta \in \Theta$ : Set  $\theta = (\theta_x)_{x \in \mathcal{S}}$  with  
 $\theta_x = (\mu_{x,1}, \dots, \mu_{x,x-1}, \mu_{x,x+1}, \dots, \mu_{x,K}, p_x)$
- $X$  follows **extended strongly coupled random walk** process with true parameter  $\theta_0$  and  $n$  firms.

## Parameter estimation:

Given observations of sample paths of  $X$ , which parameter  $\hat{\theta}$  is likeliest to be the true parameter  $\theta_0$ ?

## Likelihood function:

$$\mathcal{L}(\theta) = \left( \prod_{\substack{x,y \in S \\ x \neq y}} \prod_{\substack{a,b=1 \\ a \geq b}}^n \left( \mu_{xy} p_x^{b-1} (1 - p_x)^{a-b} \right)^{N_{x,y,a,b}} \right) \\ \times \left( \prod_{j=1}^k \mathbf{P}[X_0 = z_j] \right) \exp \left\{ - \sum_{x \in S} \mu_x \sum_{a=1}^n T_{x,a} \sum_{j=0}^{a-1} (1 - p_x)^j \right\}$$

- $N_{x,y,a,b} \in \mathbb{N}_0$ : total number of simultaneous rating changes of  $b$  firms from  $x$  to  $y \neq x$ ,  $a$  firms originally with rating  $x$  in the observed  $k$  paths
- $T_{x,a} \in [0, T]$ : total time that exactly  $a$  firms have rating  $x$
- $z_j \in S^n$ : initial rating in the  $j$ -th observed path

## Theorem (Maximum Likelihood Estimator)

The parameters in  $\hat{\Theta} \subset \Theta$  are exactly the MLE, where for  $\hat{\theta} \in \hat{\Theta}$  holds:

- 1  $\hat{p}_x$  for  $x \in S$  is either 0, 1, the unique root in  $(0, 1)$  of polynomial  $P_x$ , or arbitrary depending on  $\tilde{N}_{x,a,b}$  and  $T_{x,a}$ .
- 2 For  $x \in S$  with  $T_{x,a} > 0$  for  $a \geq 1$ :

$$\hat{\mu}_{xy} = \frac{\sum_{a=1}^n \sum_{b=1}^a N_{x,y,a,b}}{\sum_{a=1}^n T_{x,a} \sum_{j=0}^{a-1} (1 - \hat{p}_x)^j}, \quad \text{for } y \in S \text{ with } x \neq y$$

- 3  $\hat{\theta}_x$  arbitrary for  $x \in S$  with  $T_{x,a} = 0$  for all  $a \in \mathbb{N}$ .

- $\tilde{N}_{x,a,b} \in \mathbb{N}_0$ : total number of rating changes of  $b$  firms with rating  $x$ , where  $a$  firms originally in class  $x$

Model

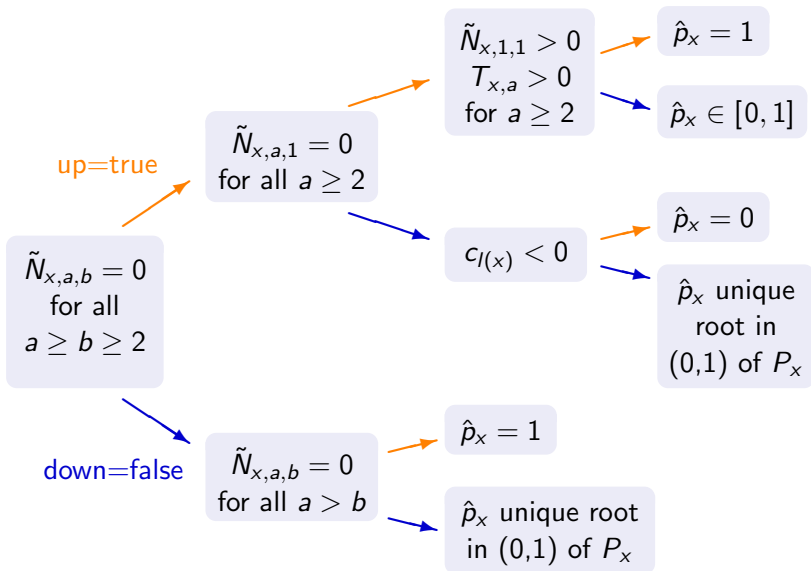
Simulation

Likelihood  
Estimation

MLE

Asymptotic  
Properties

# Maximum Likelihood Estimator for $p$



# Definition of Polynomial $P_x$

## Definition of polynomial $P_x$ :

- $P_x(p) = c_0 + c_1p + c_2p^2 + \dots + c_np^n$  has coefficients

$$c_0 = \sum_{\substack{a,b=1 \\ a \geq b}}^n (b-1) \tilde{N}_{x,a,b} \sum_{i=1}^n i T_{x,i}$$

$$c_j = (-1)^j \sum_{\substack{a,b=1 \\ a \geq b}}^n \tilde{N}_{x,a,b} \sum_{i=j}^n \binom{i}{j} \left( \frac{i-j}{j+1} b + a - i \right) T_{x,i}$$

for  $j \in \{1, \dots, n\}$

## Definition of $c_{l(x)}$ :

- $l(x) \in \{0, \dots, n\}$  is the maximal index such that  $c_j = 0$  for all  $j \in \{0, \dots, l(x) - 1\}$ .

## Theorem (Consistency)

- $\Theta \ni \theta_0 = (\mu_{x,1}, \dots, \mu_{x,x-1}, \mu_{x,x+1}, \dots, \mu_{x,K}, p_x)_{x \in S}$
- $\hat{\theta}^k$  for  $k \in \mathbb{N}$ : MLE for the observed first  $k$  paths

Assume  $\mu_x > 0$  for all  $x \in S$  and the expected time is positive, that more than one firm has rating  $x$  in the path of  $X$ .

Then the maximum likelihood estimator of  $\theta_0$  is **strongly asymptotically consistent**, i. e.

$$\hat{\theta}^k \rightarrow \theta_0, \quad \text{a. s. for } k \rightarrow \infty.$$

## Theorem (Asymptotic normality)

- $\Theta \ni \theta_0 = (\mu_{x,1}, \dots, \mu_{x,x-1}, \mu_{x,x+1}, \dots, \mu_{x,K}, p_x)_{x \in S}$
- $\hat{\theta}^k$  for  $k \in \mathbb{N}$ : MLE for the observed first  $k$  paths

Assume  $\mu_{xy} \in (0, \infty)$  and  $p_x \in (0, 1)$  for  $x, y \in S$  with  $x \neq y$ .

Then the maximum likelihood estimator is **asymptotically normal**, i.e.  $\sqrt{k}(\hat{\theta}^k - \theta_0)$  converges to a normal distribution for  $k \rightarrow \infty$  with mean zero and the inverse Fisher matrix as covariance matrix.